

Climate event attribution using multivariate peaks-over-thresholds modelling

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Outline

- 1 Introduction
- 2 Peaks-over-threshold modelling
- 3 Modelling causation probabilities
- 4 Application: European precipitation

Climate event attribution



- Environmental extreme events: increase in frequency and magnitude in view of climate change.
- Climate scientists want to measure the impact of humans on changes in the earth's climate system
→ **event attribution**.
- Use climate models to compare probabilities of an extreme event with and without **anthropogenic forcing**.

By how much have human activities increased the risk of occurrence of an extreme event?

- Compare probability of an extreme event in the factual world (p_1) and in a counterfactual world (p_0).
 - ▶ Counterfactual world: without anthropogenic forcing, i.e., without human influence.
 - ▶ Extreme event: a random variable of interest (rainfall, temperature, ...) exceeds some high threshold.

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- Many attribution studies are based on the normal distribution (not heavy-tailed), or use nonparametric estimation (no extrapolation possible), or are computationally heavy.
- We use **multivariate peaks-over-thresholds modelling** (K., Rootzén, Segers, and Wadsworth, 2018) to characterize p_0 and p_1 .

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- Under certain conditions (Hannart et al., 2016), these probabilities can be calculated by

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- The FAR coincides with the probability that human activities are necessary for the event to occur.
- When this threshold is high, the event is rare in the factual world (p_1 small) and nearly impossible in the counterfactual worlds ($p_0 \approx 0$).

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Univariate peaks over thresholds modelling

- The distribution function H of a **generalized Pareto distribution** is

$$H(x; \sigma, \gamma) = 1 - \left(1 + \frac{\gamma x}{\sigma}\right)_+^{-1/\gamma},$$

where $\sigma > 0$ and $\gamma \in \mathbb{R}$ are the scale and shape parameters.

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- If Y is a random variable and u a sufficiently high threshold,

$$Y - u \mid Y > u \stackrel{d}{\approx} X,$$

where $X \sim \text{GPD}(\sigma, \gamma)$ (Pickands-Balkema-de Haan theorem).

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- The parameter γ characterizes the heaviness of the tail and the support of the distribution.
- The GPD can be used to model the tail of Y , since for $y > u$,

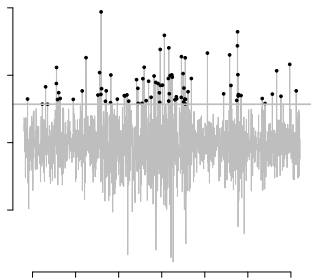
$$\mathbb{P}[Y > y] \approx \mathbb{P}[Y > u] \bar{H}(y - u; \sigma, \gamma).$$

Univariate vs multivariate peaks-over-thresholds

Univariate modelling strategy:

- Fix some high threshold u .
- Fit a GPD to the conditional threshold excesses

$$Y - u \mid Y > u.$$

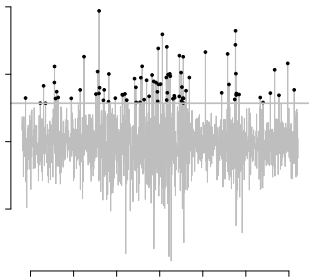


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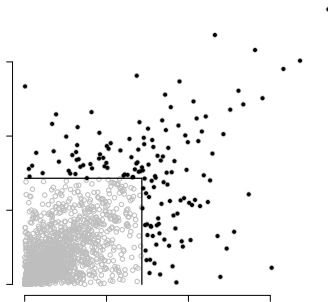
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Multivariate modelling strategy:

- Fix some high threshold u .
- Fit a multivariate GPD to the conditional threshold excesses

$$\mathbf{Y} - \mathbf{u} \mid \mathbf{Y} \not\leq \mathbf{u}.$$



Characterizing a multivariate GPD

- Let \mathbf{Y} be a d -dimensional random vector and $\mathbf{u} \in \mathbb{R}^d$ a threshold.
- If $\mathbf{X} \approx \mathbf{Y} - \mathbf{u} \mid \mathbf{Y} \not\leq \mathbf{u}$, then \mathbf{X} follows approximately a multivariate GPD (Rootzén and Tajvidi, 2006).
- The distribution function H of a multivariate GPD vector \mathbf{X} belongs to an infinite-dimensional parametric family: a model construction tool has been proposed in Rootzén et al. (2018).

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- The margins of a multivariate GPD are generally **not** GPDs, but the conditional margins are GPDs, i.e.,

$$\mathbb{P}[X_j \leq x \mid X_j > 0] = H(x; \sigma_j, \gamma_j), \quad j \in \{1, \dots, d\}.$$

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- Suppose that $\gamma := \gamma_1 = \dots = \gamma_d$, i.e., the shape parameters of the conditional margins are equal. For weights $\mathbf{w} = (w_1, \dots, w_d) > \mathbf{0}$,

$$\mathbf{w}^T \mathbf{X} \mid \mathbf{w}^T \mathbf{X} > 0 \sim \text{GPD}(\mathbf{w}^T \boldsymbol{\sigma}, \gamma).$$

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Modelling causation probabilities: univariate case

- Let $Y^{(0)}$, $Y^{(1)}$ denote the climate model output in the counterfactual and factual world respectively.
- Assume $Y^{(i)} - u^{(i)} \mid Y^{(i)} > u^{(i)} \approx X^{(i)} \sim \text{GPD}(\sigma^{(i)}, \gamma^{(i)})$ for $i = 0, 1$.

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- Set

$$\text{FAR}(v) = \text{PN}(v) = 1 - \frac{p_0(v)}{p_1(v)},$$

where $p_i(v) = \mathbb{P}[Y^{(i)} > v]$.

- For $v > u^{(i)}$,

$$p_i(v) \approx \mathbb{P}\left[Y^{(i)} > u^{(i)}\right] \overline{H}\left(v - u^{(i)}; \sigma^{(i)}, \gamma^{(i)}\right).$$

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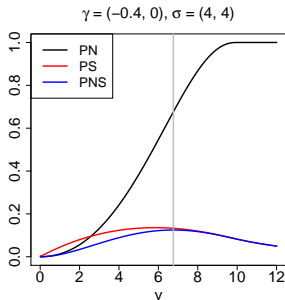
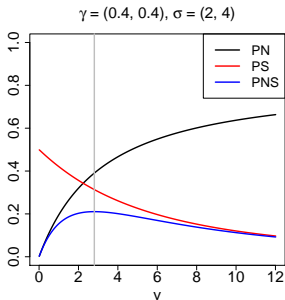
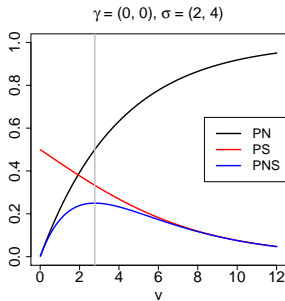
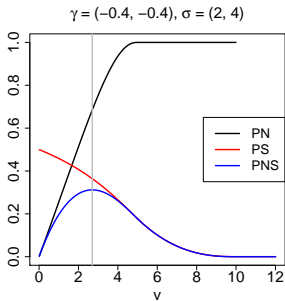
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- The probability $\mathbb{P}[Y^{(i)} > u^{(i)}]$ can be estimated by the empirical excess probability of $u^{(i)}$.
- Estimating p_0 and p_1 for high values of v then amounts to estimating the parameters of a univariate GPD.

The PN, PS and PNS for different GPD parameters



Modelling causation probabilities: multivariate case

- Let $\mathbf{Y}^{(0)}, \mathbf{Y}^{(1)} \in \mathbb{R}^d$ be random vectors of climate model output in the counterfactual and factual world respectively.
- Assume $\mathbf{Y}^{(i)} - \mathbf{u}^{(i)} \mid \mathbf{Y}^{(i)} \not\leq \mathbf{u}^{(i)} \approx \mathbf{X}^{(i)}$ for some high threshold $\mathbf{u}^{(i)}$, where $\mathbf{X}^{(i)}$ is a multivariate GPD vector.

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- Set

$$\text{PN}(v; \mathbf{w}) = \text{FAR}(v; \mathbf{w}) = 1 - \frac{p_0(v; \mathbf{w})}{p_1(v; \mathbf{w})},$$

for weights $\mathbf{w} = (w_1, \dots, w_d)^T > \mathbf{0}$ with $\sum_{j=1}^d w_j = 1$, and

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- Recall the sum-stability property,

$$\mathbb{P}[\mathbf{w}^T \mathbf{X}^{(i)} > v] = \mathbb{P}[\mathbf{w}^T \mathbf{X}^{(i)} > 0] \bar{H}(v; \mathbf{w}^T \boldsymbol{\sigma}^{(i)}, \gamma^{(i)}).$$

Modelling causation probabilities: multivariate case

- The sum-stability property is only valid when $\gamma^{(i)} := \gamma_1^{(i)} = \dots = \gamma_d^{(i)}$ for $i \in \{0, 1\}$.
- Combining $\mathbf{Y}^{(i)} - \mathbf{u}^{(i)} \mid \mathbf{Y}^{(i)} \not\leq \mathbf{u}^{(i)} \approx \mathbf{X}^{(i)}$ and sum-stability, we find

$$p_i(v) \approx \mathbb{P} \left[\mathbf{w}^T \mathbf{Y}^{(i)} > \mathbf{w}^T \mathbf{u}^{(i)} \right] \overline{H} \left(v - \mathbf{w}^T \mathbf{u}^{(i)}; \mathbf{w}^T \boldsymbol{\sigma}^{(i)}, \gamma^{(i)} \right)$$

for $v > \mathbf{w}^T \mathbf{u}^{(i)}$.

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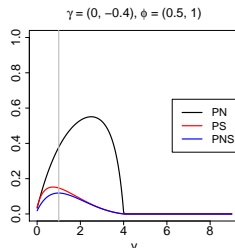
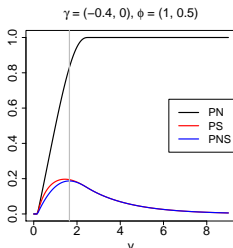
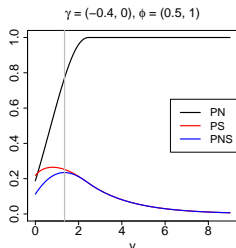
for $v > \mathbf{w}^T \mathbf{u}^{(i)}$.

- The first term on the right-hand side can be estimated non-parametrically.
- The parameters σ_j, γ can be estimated by fitting GPDs to the conditional marginal threshold exceedances.

The PN, PS and PNS for different scenarios

Examples of causation probabilities in $d = 2$ with $\mathbf{w} = (1/2, 1/2)^T$ and $\gamma = \gamma^{(0)} = \gamma^{(1)}$, where, from left to right,

- 1 The factual world exhibits **increased dependence** and **increased marginal tail heaviness**.
- 2 The factual world shows **decreased dependence** and **increased marginal tail heaviness**.
- 3 The factual world shows **increased dependence** and **decreased marginal tail heaviness**.



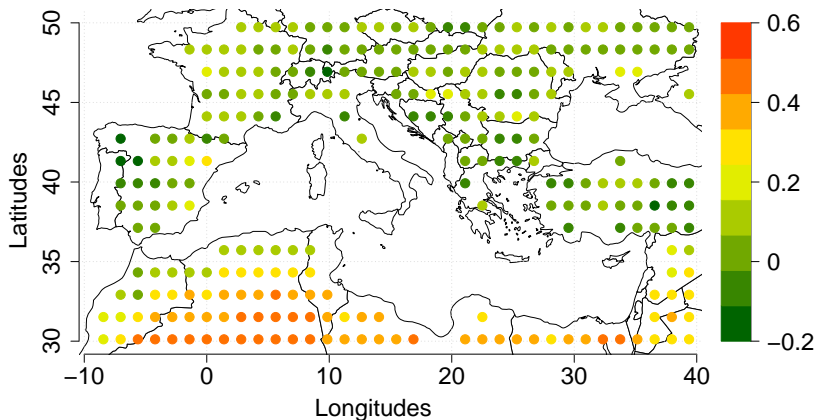
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European precipitation: shape parameters of GPD fit

Weekly maximum precipitation amounts (mm/day) between 1850 and 2005 on a grid of $d = 268$ points in central Europe.

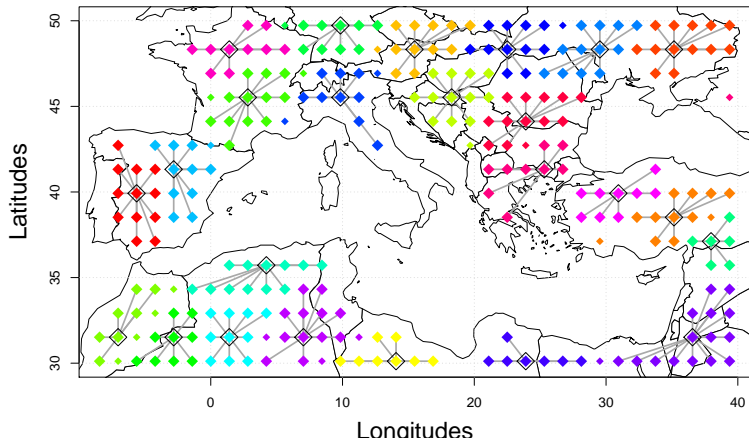
Shape parameter in the factual world



European precipitation: clusters

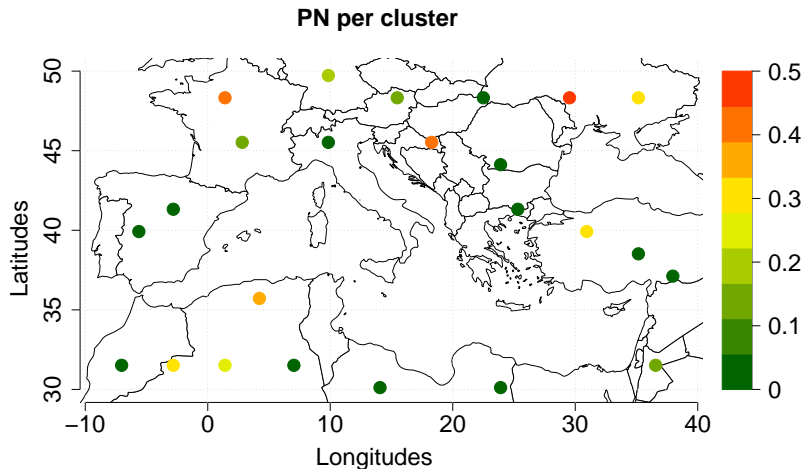
Partitioning around medoids (PAM) algorithm (Kaufman and Rousseeuw, 1990) based on madogram distance.

K = 24 clusters



European precipitation: probabilities of necessary causation

Fraction of Attributable Risk (or PN) per cluster, evaluated at a threshold v corresponding to a **hundred-year return level** of $\mathbf{w}^T \mathbf{Y}^{(1)}$.



Discussion

- For $v > \mathbf{w}^T \mathbf{u}$, we can also use the approximation

$$\mathbb{P}[\mathbf{w}^T \mathbf{Y} > v] \approx \mathbb{P}[\mathbf{Y} \not\leq \mathbf{u}] \mathbb{P}[\mathbf{w}^T \mathbf{X} > 0] \bar{H}(v - \mathbf{w}^T \mathbf{u}; \mathbf{w}^T \boldsymbol{\sigma}, \gamma).$$

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- For certain parametric models on the multivariate GPD vector \mathbf{X} and for certain γ , the probability $\mathbb{P}[\mathbf{w}^T \mathbf{X} > 0]$ can be calculated analytically.
- To enhance the FAR, we look for weights w_1, \dots, w_d that are optimal, i.e.,

$$\mathbf{w}_{\text{opt}} = \arg \max_{\mathbf{w}: \sum_j w_j = 1} \text{FAR}(v; \mathbf{w}).$$

→ analytic expression for \mathbf{w}_{opt} only available in $d = 2$ for special cases.

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