

The genesis and use of time-varying frailty models for representing heterogeneities in the transmission of infectious diseases

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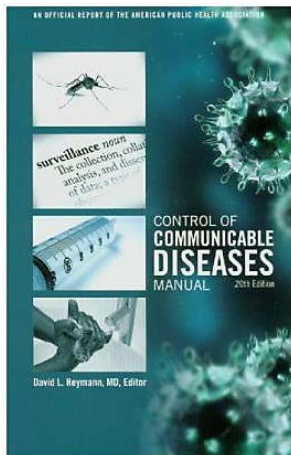
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Infectious diseases and their transmission



- **Infectious diseases** are caused by pathogenic biological agents.
- The **spreading** of infectious agents is called **transmission**.
- Example: **measles** is transmitted from person to person primarily by the **airborne route**.
- The majority of transmission models are deterministic **compartmental models**.

The Susceptible-Infected-Recovered (SIR) model

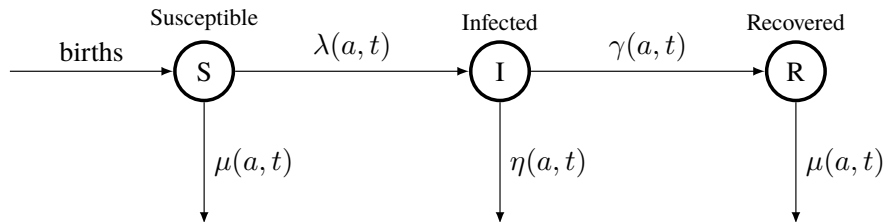


Figure: Flow diagram for the compartmental SIR model: Individuals are born into the susceptible class **S** and move to the infected state **I** at rate $\lambda(a, t)$, after which they recover and move to **R** at rate $\gamma(a, t)$. All individuals are subject to natural mortality at rate $\mu(a, t)$ and infected individuals to an additional disease-related mortality at rate $\alpha(a, t)$. It is assumed that $\eta(a, t) = \mu(a, t) + \alpha(a, t)$.

Time-homogeneous SIR model

- The **time-homogeneous SIR model** can be described using the following set of ordinary differential equations (ODEs) in age:

$$\begin{aligned}\frac{dS(a)}{da} &= -[\lambda(a) + \mu(a)] S(a) , \\ \frac{dI(a)}{da} &= \lambda(a)S(a) - [\gamma(a) + \eta(a)] I(a) , \\ \frac{dR(a)}{da} &= \gamma(a)I(a) - \mu(a)R(a) ,\end{aligned}$$

where $S(a)$, $I(a)$ and $R(a)$ represent the number of susceptible, infected and recovered individuals of age a .

- The total number of individuals of age a in the population is $N(a) = S(a) + I(a) + R(a)$.

Susceptible-Infected-Recovered-Susceptible (SIRS) model

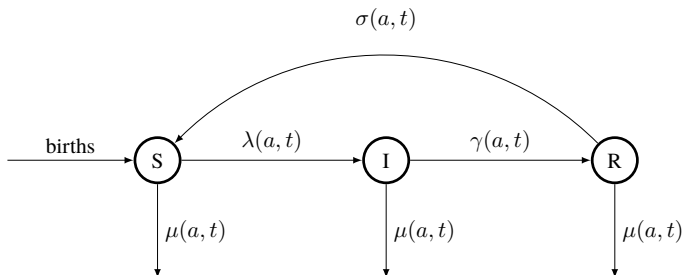


Figure: Flow diagram for the time-heterogeneous SIRS model: Individuals are born into the susceptible class **S** and move to the infected state **I** at rate $\lambda(a, t)$, after which they recover and move to **R** at rate $\gamma(a, t)$. Subsequently, individuals lose protective immunity and move back to **S** at replenishment rate $\sigma(a, t)$. All individuals are subject to natural mortality at rate $\mu(a, t)$.

Heterogeneities in the transmission of infectious diseases

- **Individuals** in a population show **variation** with respect to properties that are relevant to the **transmission of infections**.
- **Heterogeneities** exist due to variation between individuals in
 - **susceptibility** to infection;
 - **infectiousness**, once infected;
 - **activity levels** in interacting with other individuals.
- Heterogeneity of a population may affect both
 - 1 the way in which infections are transmitted within it,
 - 2 and the effectiveness of strategies to control them.
- Allowing for individual heterogeneity in **statistical** and **mathematical models** of infectious diseases is important.

How to quantify heterogeneities?

- Models often involve specifying **contact rates** between individuals.
- A contact is an event during which transmission of infection between two individuals **could** occur.
- For **most types** of infection, there is **no event** that can be clearly or uniquely defined as a contact.
- For these infections it is usually necessary to define a contact by some **proxy variable**.

Quantifying heterogeneities: frailty modelling approach

- A **different approach** to make inferences on heterogeneities uses the fact that they leave an **epidemiological footprint**.
- The **extent of heterogeneity** in behaviour relevant to the transmission of infection will be reflected by the **strength of the association** between infections.
- The degree of heterogeneity can be estimated using **multivariate frailty models** for the hazard of infection.
- This approach enables us to observe the effects of heterogeneity without explicitly specifying the mechanisms that give rise to them.

Modelling individual effects

- To facilitate the notation, it is assumed that **age**, denoted a , is the only **measured attribute** of an individual.
- Each individual has **unobserved (latent) characteristics** \mathbf{z} with density $f(\mathbf{z})$; \mathbf{z} comprises age-invariant random variables z_1, \dots, z_K .
- We suppose that the age-dependent effect of the latent characteristics can be compounded into a single random variable $w(a, \mathbf{z})$, where $w(\cdot)$ is a deterministic function.
- For each a , $w(a, \mathbf{z})$ has mean 1. Of key importance in describing the **degree of heterogeneity is the variance** of $w(a, \mathbf{z})$.

Effective contacts

- An **effective contact** is defined as an event involving individuals X and Y such that, if Y was infectious and X susceptible, then Y would infect X .
- Let $\beta(a, \mathbf{z}; a', \mathbf{z}')$ represent the per-capita rate at which an individual with characteristics (a', \mathbf{z}') makes effective contacts with individuals with characteristics (a, \mathbf{z}) .
- The function $\beta(a, \mathbf{z}; a', \mathbf{z}')$ is non-negative and determines the so-called **effective contact rate surface**.

Effective contact rate surface

- The effective contact rate surface may be written as

$$\beta(a, \mathbf{z}; a', \mathbf{z}') = \alpha(a, \mathbf{z}; a', \mathbf{z}') \beta_0(a, a') ,$$

where

$$\beta_0(a, a') = \int_{\mathbf{z}'} \int_{\mathbf{z}} \beta(a, \mathbf{z}; a', \mathbf{z}') f(\mathbf{z}) f(\mathbf{z}') d\mathbf{z} d\mathbf{z}'$$

is the **average effective contact rate**.

- We assume that $\alpha(a, \mathbf{z}; a', \mathbf{z}') = w(a, \mathbf{z}) w'(a', \mathbf{z}')$, hence

$$\beta(a, \mathbf{z}; a', \mathbf{z}') = w(a, \mathbf{z}) \beta_0(a, a') w'(a', \mathbf{z}') .$$

Genesis of time-varying frailty models

- Let $\lambda(a, \mathbf{z}, t)$ be the **hazard (or force) of infection** acting on a susceptible individual of characteristics (a, \mathbf{z}) at time t .
- When the infection is in **endemic equilibrium**, the hazard of infection is of the form

$$\lambda(a, \mathbf{z}) = \int_0^\infty \int_{\mathbf{z}'} \beta(a, \mathbf{z}; a', \mathbf{z}') I(a', \mathbf{z}') d\mathbf{z}' da' ,$$

where $I(a', \mathbf{z}')$ is the number of infectious individuals with characteristics (a', \mathbf{z}') .

- The functional form which is taken by $I(a', \mathbf{z}')$ depends on whether the infection is SIR, SIRS, or some other type.

Genesis of time-varying frailty models

- The integral equation can be written as

$$\begin{aligned}
 \lambda(a, \mathbf{z}) &= \int_0^\infty \int_{\mathbf{z}'} w(a, \mathbf{z}) \beta_0(a, a') w'(a', \mathbf{z}') I(a', \mathbf{z}') \, d\mathbf{z}' \, da' \\
 &= w(a, \mathbf{z}) \underbrace{\int_0^\infty \int_{\mathbf{z}'} \beta_0(a, a') w'(a', \mathbf{z}') I(a', \mathbf{z}') \, d\mathbf{z}' \, da'}_{=\lambda_0(a)} \\
 &= w(a, \mathbf{z}) \times \lambda_0(a) \, ,
 \end{aligned}$$

where $\lambda_0(a)$ is the baseline force of infection.

- The equation above defines an **age-varying frailty model** for the hazard of infection with age-dependent frailty $w(a, \mathbf{z})$.

Time-varying frailty models

Bivariate setting:

- Consider **two infections**. For infection j the force of infection at age a for an individual with **age-varying frailty** $z_j(a)$ is assumed to be of the form

$$\lambda_j(a, z_j(a)) = z_j(a)\lambda_{0j}(a) \quad \text{for } j = 1, 2,$$

where $\lambda_{0j}(a)$ are the baseline hazards.

- We still need to...
 - find a function $w(a, z_{j1}, z_{j2}, \dots, z_{jK}) = z_j(a)$, where z_{jk} ($j = 1, 2; k = 1, \dots, K$) are **independent age-invariant frailties**;
 - make a decision whether to use **shared frailties** with $z(a) := z_1(a) = z_2(a)$ or **correlated frailties**.

Piecewise-constant frailties

- One could build **piecewise-constant frailty models** on disjoint age intervals $I_k = (a_{k-1}, a_k]$ for $k = 1, \dots, K$ with $a_0 = 0$ and $a_K < \infty$.

- Let

$$z_j(a) = \sum_{k=1}^K z_{jk} I_k(a) ,$$

where $z_{jk} > 0$ are identically distributed with unit mean and variance γ_{jk} ($j = 1, 2; k = 1, \dots, K$), and $I_k(a) = 1$ if $a \in I_k$ (with $I_k(a) = 0$ otherwise).

- Assumption: the frailty in age group k is **independent** from the frailty in age group $k + 1$.

Multiplicative family

- Consider the **multiplicative family of models**:

$$z_j(a) = \prod_{k=1}^K [1 + (z_{jk} - 1) h_{jk}(a)] \quad , \quad 0 \leq h_{jk}(a) \leq 1 \quad ,$$

where z_{jk} for $j = 1, 2$ and $k = 1, \dots, K$ are independent random variables with unit mean and variance γ_{jk} .

- The $h_{jk}(a)$ are deterministic functions such as

$$h_{jk}(a) = \exp \left[- (a\phi_{jk})^2 \right] \quad ,$$

where $\phi_{jk} > 0$ is an exponential decay parameter.

- Assumption: the frailties across age groups are perfectly correlated.

One-component time-varying shared frailty model

- Suppose that $K = 1$ and that the frailty components z_{j1} ($j = 1, 2$) follow a gamma distribution, denoted $\Gamma(\cdot, \cdot)$.
- In the shared frailty model, the correlation between the frailty terms z_{11} and z_{21} is unity we define $z_1 := z_{11} = z_{21}$.
- A one-component **age-varying shared gamma frailty model** is then

$$z(a) = [1 + (z_1 - 1)h_1(a)] ,$$

where $z_1 \sim \Gamma(\gamma_1^{-1}, \gamma_1^{-1})$.

One-component time-varying correlated frailty model

- The correlated frailty model allows for a more flexible correlation structure among the frailty terms.
- One can build a one-component **age-varying correlated gamma frailty model** as follows:

$$z_j(a) = [1 + (z_{j1} - 1)h_{j1}(a)] ,$$

where $z_{j1} = \gamma_{j1}(y_{01} + y_{j1})$, $y_{l1} \sim \Gamma(k_{l1}, 1)$ and $\gamma_{j1} = (k_{01} + k_{j1})^{-1}$ ($j = 1, 2; l = 0, 1, 2$).

- The implied correlation between between the frailty terms z_{11} and z_{21} is

$$\rho = \frac{k_{01}}{\sqrt{(k_{01} + k_{11})(k_{01} + k_{21})}} , \quad 0 \leq \rho \leq \min \left\{ \sqrt{\frac{\gamma_{11}}{\gamma_{21}}}, \sqrt{\frac{\gamma_{21}}{\gamma_{11}}} \right\} .$$

Bivariate serological survey data

- T_1 and T_2 : ages at the onset of infection by two distinct infectious agents.
- Association between T_1 and T_2 can be examined using paired serological survey data on two infections.
- Data are obtained by testing blood serum residues for the presence of antibodies to one or more infections.
- A positive (negative) results indicates prior infection (lack of prior infection), giving rise to current status data.

Observable data

- For current status data, only the information about whether the survival time of interest lies **before or after** the monitoring time (age) a is available.
- Observed information in a bivariate setting is $\{a, \delta_1, \delta_2\}$, where

$$\delta_j = \begin{cases} 1 & \text{if } T_j \leq a, \\ 0 & \text{if } T_j > a, \end{cases} \quad (j = 1, 2) \quad .$$

- Aggregated data at each age a : $(n_{00a}, n_{01a}, n_{10a}, n_{11a})$ and $n_a = \sum_{i,j=0,1} n_{ija}$.

Estimation

- Given **parameterizations** of $w(a, \mathbf{z})$, $\lambda_{01}(a)$ and $\lambda_{02}(a)$, the model is fitted by maximizing a **multinomial likelihood**.
- The multinomial log-likelihood kernel is

$$\sum_a \sum_{i,j=0,1} n_{ija} \ln \{p_{ij}(a)\} \quad ,$$

where the probabilities $p_{ij}(a)$ (in an SIR setting) are computed as

$$p_{00}(a) = E \left(\exp \left\{ - \int_0^a w(y, \mathbf{z}) [\lambda_{01}(y) + \lambda_{02}(y)] dy \right\} \right) \quad ,$$

$$p_{01}(a) = E \left(\exp \left\{ - \int_0^a w(y, \mathbf{z}) \lambda_{01}(y) dy \right\} \right) - p_{00}(a) \quad ,$$

$$p_{10}(a) = E \left(\exp \left\{ - \int_0^a w(y, \mathbf{z}) \lambda_{02}(y) dy \right\} \right) - p_{00}(a) \quad ,$$

$$p_{11}(a) = 1 - p_{01}(a) - p_{10}(a) - p_{00}(a) \quad .$$

Fitting procedure for a pre-specified model

- For the current set of parameters,
 - ① obtain the baseline hazards λ_{0j} ($j = 1, 2$),
 - ② compute the probabilities $p_{00}(a)$, $p_{01}(a)$, $p_{10}(a)$ and $p_{11}(a)$,
 - ③ evaluate the log-likelihood,and iterate until convergence.
- Possible **parameterizations of the baseline hazards** include continuous parametric baselines (such as the Gompertz hazard) or piecewise constant baselines.
- For some of the models, the expressions $p_{ij}(a)$ for $i, j = 0, 1$ **cannot** be computed in closed-form.

Applications

Description of the data

- ① Hepatitis A virus (HAV) and hepatitis B virus (HBV) serology
 - **Different** transmission routes.
 - Data obtained from a seroepidemiological study undertaken in 1993 and 1994 in Flanders, Belgium. In total, 4026 blood samples were drawn.
- ② Parvovirus B19 and varizella zoster virus (VZV) serology
 - **Similar** transmission routes.
 - Data for 3379 individuals between 2001 and 2003 in Belgium.
 - Parvovirus B19: immunizing process (SIR) or recurrent infection process (SIRS).
 - Estimation of the **basic reproduction number**, R_0 , from serological data and **social contact data**.
 - Social contact hypothesis: $\beta_0(a, a') = q \times c(a, a')$, where q is an infection-specific proportionality factor.

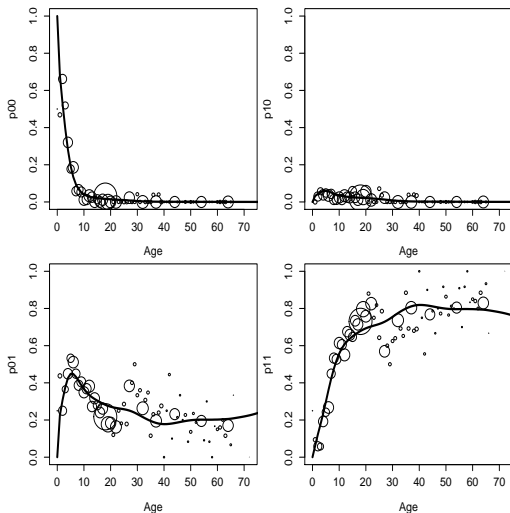
Fitting results for HAV and HBV infection data

Frailty model	Frailty parameters	Estimates (s.e.)	AIC	BIC
SGF	$\sqrt{\gamma_1}$	0.725 (0.086)	5824.90	5856.41
CGF	$\sqrt{\gamma_{11}}$	1.651 (0.176)	5794.89	5828.99
	$\sqrt{\gamma_{21}}$	1.608 (2.272)		
	ρ	0.497 (0.702)		
ADSGF-1C	$\sqrt{\gamma_1}$	5.843 (0.829)	5756.01	5793.82
	ϕ	0.034 (0.005)		
ADCGF-1C	$\sqrt{\gamma_{11}}$	6.606 (1.020)	5757.04	5807.44
	$\sqrt{\gamma_{21}}$	5.765 (0.831)		
	ϕ	0.025 (0.007)		
	ρ	0.871 (0.080)		
ADSGF-2C	$\sqrt{\gamma_1}$	5.814 (0.446)	5758.03	5802.13
	$\sqrt{\gamma_2}$	0.009 (0.124)		
	ϕ	0.034 (0.005)		
ADPiecewiseSGF	$\sqrt{\gamma_1}$	3.671 (0.606)	5749.01	5799.42
	$\sqrt{\gamma_2}$	2.421 (0.504)		
	$\sqrt{\gamma_3}$	0.012 (0.160)		
	$\sqrt{\gamma_4}$	8.813 (7.856)		

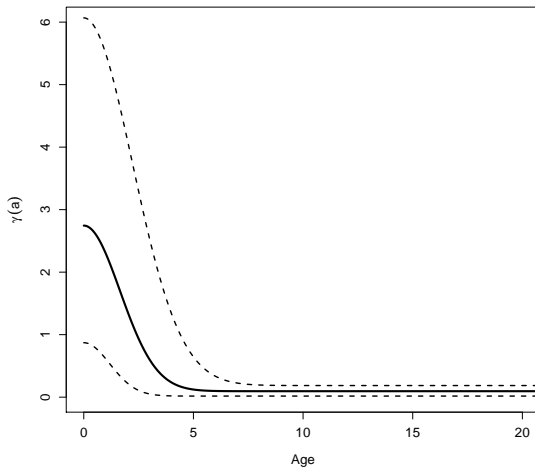
Fitting results for Parvovirus B19 and VZV infection data

Frailty model	Parameters	Estimates [CI]	\hat{R}_0 [CI]	AIC	BIC
SGF-SIR	q_1	0.072 [0.069, 0.075]	3.60 [3.35, 3.88]	4937.14	4955.51
	q_2	0.200 [0.188, 0.214]	11.64 [10.59, 12.82]		
	γ	0.152 [0.118, 0.188]			
ADSGF-1C-SIR	q_1	0.072 [0.069, 0.076]	3.60 [3.22, 3.99]	4939.14	4963.64
	q_2	0.200 [0.183, 0.221]	11.64 [9.99, 13.49]		
	γ	0.152 [0.100, 0.210]			
ADSGF-2C-SIR	ϕ	0.000 [0.000, 0.009]		4912.08	4942.70
	q_1	0.066 [0.062, 0.071]	3.74 [3.15, 4.87]		
	q_2	0.235 [0.191, 0.299]	15.65 [11.38, 24.08]		
	γ_1	2.918 [1.524, 5.004]			
	γ_2	0.233 [0.156, 0.323]			
	ϕ	0.316 [0.246, 0.425]			
SGF-SIRS	q_1	0.071 [0.068, 0.074]	3.18 [2.97, 3.43]	4869.83	4894.33
	σ	0.011 [0.008, 0.015]			
	q_2	0.173 [0.163, 0.183]	8.98 [8.22, 9.83]		
	γ	0.032 [0.002, 0.065]			
ADSGF-1C-SIRS	q_1	0.065 [0.061, 0.070]	2.90 [2.64, 3.49]	4862.93	4893.56
	σ	0.012 [0.009, 0.016]			
	q_2	0.158 [0.141, 0.179]	8.19 [7.15, 10.46]		
	γ	1.470 [0.415, 3.498]			
ADSGF-2C-SIRS	ϕ	0.330 [0.209, 0.530]		4859.26	4896.01
	q_1	0.066 [0.062, 0.071]	3.30 [2.79, 4.45]		
	σ	0.011 [0.007, 0.015]			
	q_2	0.193 [0.156, 0.257]	11.27 [8.11, 18.90]		
	γ_1	2.419 [0.839, 4.960]			
	γ_2	0.095 [0.017, 0.186]			
	ϕ	0.303 [0.226, 0.423]			

Observed and fitted seroprevalence of B19 and VZV



Age-varying shared frailty variance



Concluding remarks

- Time-varying frailty models are a natural choice for capturing individual heterogeneities relevant to the transmission of infectious diseases.
- Multivariate frailty models with shared/correlated frailties can be used for
 - inducing association between infection times within individuals,
 - heterogeneity among individuals.
- Central to our approach is the use of paired serological survey data on different infections for the same individuals.
- For pairs of infection with the same transmission route, a shared frailty model seems appropriate.
- Frailty modelling is fraught with **lack of identifiability**.
- Further work, some of it under way, is required in several areas.

Research grant

- 3-year grant entitled “Frailty modelling for multivariate current status data with applications in epidemiology”



- Research project funded by the
- Aims:
 - 1 to develop innovative statistical approaches to analyse multivariate current status data,
 - 2 to develop estimation methods for the new models,
 - 3 to provide statistical software and examples of applications for the new methodologies.
- External collaborator: Niel Hens
- I am currently seeking a promising PhD student (or Postdoc) to work on this project.

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